

A Study of Causality between Bank Credit and GDP Growth in India

–Dhiren Jotwani*

ABSTRACT

The factors that cause Economic growth are varied and diverse. Theory has however managed to identify certain key factors, of which finance is one. In recent years, there have been studies using econometric time-series analysis to study the short-run and long-run relationships between finance and growth, for various countries. This paper is a study of the Indian economy, where we try to determine the causal relationship between bank credit and economic growth. Data from 1972 to 2012 for the Indian economy has been used for this study. The results suggest that provision of bank credit leads to economic growth. However, an increase in economic growth may not lead to further provision of bank credit in the economy. In other words, there is unidirectional causality from bank credit to growth.

Keywords: *Economic Growth, Causality, Time-Series Analysis, Bank Credit*

INTRODUCTION

An inquiry of what causes Economies to grow at a certain pace, their differences in growth patterns, and their long run paths of growth, is perhaps one of the oldest and yet still most widely pursued questions in Economics. This question has a direct implication on business administration and business policy. Over the years, scholars have used various approaches, from discussions to complex mathematics, to study this phenomenon.

Economic growth in any economy is achieved as a result of various factors. The key factors are those that lead to capital accumulation and technological progress. Human capital and physical capital are essential

* Assistant Professor, Economics and Finance, Nirma University, Ahmedabad, Gujarat, India. E-mail: dhirenjotwani@hotmail.com

in the process. The private and public sector needs to make investments in capital for growth to occur. This is where finance is so crucial. For any kind of investment to occur, sources of finance are necessary. Similarly, for improvements in technology to occur, finance is necessary.

In India, the financial sector is very large and widespread, comprising of various entities. However, the main sources of finance for corporations are still banks and stock markets. Further, for regular and immediate needs of corporations, banks are very often the preferred means of raising finance. Hence, a study of how bank credit determines overall growth in the economy is quite interesting.

LITERATURE REVIEW

Current reviews suggest that Bagehot (1873) is one of the earliest literary works that links financial development and economic growth. His study gave illustrations of how money market developments in England aided the flow of capital towards productive industries. The study also articulated the relation between finance, trade and growth. Further, Schumpeter (1912) described how financial intermediation is the centre of economic development. Several other prominent economists have studied this relation and made immense theoretical and empirical contributions [Patrick (1966), Goldsmith (1969), McKinnon (1973) and Shaw (1973)].

Over the years, new theories of economic growth were developed. The most recent theory of endogenous growth very appropriately blends the concept of financial development with economic growth. The prominent initial contributions in this area were by Greenwood and Jovanovic (1990), Pagano (1993) and King and Levine (1993). Their research shows that financial development does have a positive impact on economic growth through investment, saving, productivity of capital and effective management of information.

Further, there is no particular differentiation needed between the proportion of banks or stock markets in the economy. Although this argument still goes on, it is believed that both banks and stock markets behave as complementary, rather than rivals in the finance-growth nexus. This is shown by Boyd and Prescott (1986), Boyd and Smith (1998) and Blackburn and Capasso (2005).

The techniques used for econometric analysis have also evolved over the years. From basic descriptive analyses, to cross sectional studies,

the focus has now gradually shifted to multivariate models. In recent years, researchers have used techniques of time-series analysis. The most prominent ones are Jung (1986), Demetriades and Hussein (1996), Neusser and Kugler (1998), Bell and Rousseau (2001), Rousseau and Sylla (2005), among others.

Among earlier and recent work in the Indian context, some important ones are Acharya et al (2009), Bell and Rousseau (2001), Chakraborty (2007, 2010, 2011), Demetriades and Luintel (1996), Mishra et al (2010), Oura (2008) and Pradhan (2009, 2010a, 2010b, 2011).

ECONOMETRIC TECHNIQUE AND RESULTS

Continuing with the preceding discussion, current studies on the finance-growth nexus try to analyze the extent and nature of causality between the two variables. While it is well established that finance indeed plays an important role in economic growth, there are debates over the issue of causality. For this, the usage of certain econometric techniques is required. The data used in this study is time-series data, from 1972 to 2012. It has been collected from the extensive database of the Reserve Bank of India.

Many current studies have employed cointegration and granger causality tests. Time-series analysis of data requires various diagnostic checks. For non time-series data, we can directly proceed to our data analysis technique. However, for time-series data, the very first check is that of stationarity. The Augmented Dickey-Fuller test is employed here. The results were also checked with other tests, and confirmed to be consistent.

Table 1: Tests of Stationarity

Name of Variable	DF Statistic at Level		DF Statistic at First Difference	
	Constant	Constant and Trend	Constant	Constant and Trend
Total GDP Growth (TGG)	0.20	-2.24	-3.32***	-4.08***
Total Bank Credit (TBC)	0.74	-1.22	-3.81***	-4.24***
Bank Credit to Industry (BCI)	3.05	-0.89	-6.76***	-7.32***

***, **, * indicates DF test value is significant at 1%, 5% and 10% level of significance respectively.

For constant model, critical values at 1%, 5% and 10% level of significance are, -2.62, -1.95 and -1.61 respectively

For constant and trend model, critical values at 1%, 5% and 10% level of significance are -3.77, -3.19 and -2.89 respectively

Name of Variable	ADF Statistic at Level		ADF Statistic at First Difference	
	Constant	Constant and Trend	Constant	Constant and Trend
Total GDP Growth (TGG)	0.19	-2.29	-4.53***	-4.48***
Total Bank Credit (TBC)	2.41	0.20	-3.83***	-4.15**
Bank Credit to Industry (BCI)	1.58	-0.37	-6.70***	-7.23***

***, **, * indicates ADF test value is significant at 1%, 5% and 10% level of significance respectively.

For constant model, critical values at 1%, 5% and 10% level of significance are, -3.61, -2.94 and -2.61 respectively

For constant and trend model, critical values at 1%, 5% and 10% level of significance are -4.21, -3.53 and -3.19 respectively

For the sake of brevity, we have named Total GDP Growth as TGG, Total Bank Credit as TBC and Bank Credit to Industry as BCI for the remainder of the text.

The tests of stationarity have to be performed for different possibilities, all of which have been included in the table. All three of our variables are not stationary at levels. This implies that we may need to test first differences for the same. If we run our analysis on level data, we are most likely to get meaningless or spurious results. Thus, we further perform the tests of stationarity at first differences. Now, all the calculated ADF values exceed the critical values. These indicate that all three variables are stationary at first differences and we can proceed to pair wise cointegration tests. We first begin with the cointegration between TGG and TBC.

Cointegration between TGG and TBC

Selecting the appropriate lag length is crucial, as the test of cointegration is sensitive to lag lengths of the VAR system. We select the lag length that fulfils various criteria. In this case, it is 2. This is shown in table 2 below.

Table 2: Lag Length Selection

Lag	LogL	LR	FPE	AIC	SC	HQ
0	-63.26	NA	0.11	3.43	3.52	3.46
1	135.12	365.46	0.00	-6.79	-6.54*	-6.70
2	142.29	12.45*	0.00*	-6.96*	-6.53	-6.81*
3	144.59	3.75	0.00	-6.87	-6.27	-6.66

* indicates lag order selected by the criterion; LR: sequential modified LR test statistic (each test at 5% level); FPE: Final prediction error; AIC: Akaike information criterion; SC: Schwarz information criterion; HQ: Hannan-Quinn information criterion

Table 3: Statistics to Check for Cointegration

Null	Alternate	Trace Statistic	Max-eigen Value Statistic
$r=0$	$r=1$	28.2 (20.2)***	25.97 (15.89)***
$r \leq 1$	$r=2$	2.23 (9.16)	2.23 (9.16)

Figure in parenthesis is critical value; *** denotes rejection of hypothesis at 0.01 level, ** denotes rejection at 0.05 level

The results indicate that there exists one cointegrating vector. Both the criteria, i.e., the trace statistic and the max-eigen value statistic, indicate the presence of one cointegrating vector, and the pantula principle selects model 2, i.e., intercept in CE and no trend or intercept in VAR. This model selection is once again based on minimizing of certain criteria.

Table 4: Cointegrating Vector

	TGG	TBC	C
Coint. Vector	1.00	-1.85	5.67

Normalized ECM: $TGG = 1.85TBC - 5.67$

There is one cointegrating vector, and the TBC shows positive sign. This indicates a positive relationship between TBC and TGG.

This table gives the estimated coefficients of the variables in the ECM. We find that TGG has a statistically significant error correction coefficient with the correct sign. The adjustment coefficients indicate that 4.3% disequilibrium in TGG is corrected each year. However, no significant error correction is seen in TBC. This makes it weakly exogenous. But, t-values for TBC are significant. Therefore, in the long run, there seems to be unidirectional causality from TBC to TGG. This is indicated by the individual ECM regressions as well.

Table 5: Adjustment Coefficients

Error Correction	D(TGG)	D(TBC)
CointEq1	-0.044	0.001
	[-5.44]***	[0.073]
D(TGG(-1))	0.113	0.132
	[0.68]	[0.376]
D(TGG(-2))	-0.279	0.236
	[-1.963]**	[0.788]
D(TBC(-1))	-0.296344	0.404236
	[-3.21331]***	[2.07986]**
D(TBC(-2))	-0.134919	0.163561
	[-1.28352]	[0.73833]
R-squared	0.40	0.203
F-statistic	5.508	2.105

Figures in parentheses are t-values; *** denotes significance at 0.01 level; ** at 0.05 level; * at 0.10 level

Cointegration between TGG and BCI

Once again, selection of the appropriate VAR lag length is the first crucial step in the test of cointegration. We select the lag length that fulfils various criteria. We see that, once again, it is 2. This is shown in table 6 below.

Table 6: Lag Length Selection

Lag	LogL	LR	FPE	AIC	SC	HQ
0	-56.19	NA	0.07	3.06	3.15	3.09
1	115.56	316.39	0.00	-5.77	-5.51	-5.67
2	123.15	13.18*	0.00*	-5.96*	-5.52*	-5.80*
3	124.64	2.43	0.00	-5.82	-5.22	-5.61

* indicates lag order selected by the criterion; LR: sequential modified LR test statistic (each test at 5% level); FPE: Final prediction error; AIC: Akaike information criterion; SC: Schwarz information criterion; HQ: Hannan-Quinn information criterion

Next, we once again apply the pantula principle to test for cointegration.

The results indicate that there exists one cointegrating vector. Both the criteria, i.e., the trace statistic and the max-eigen value statistic, indicate the presence of one cointegrating vector, and the pantula principle selects model 1, i.e., no intercept or trend or intercept in CE and VAR. This model selection is once again based on minimizing of certain criteria.

Table 7: Statistics to Check for Cointegration

Null	Alternate	Trace Statistic	Max-eigen Value Statistic
$r=0$	$r=1$	18.97 (12.32)***	15.74 (11.22)***
$r \leq 1$	$r=2$	3.22 (4.13)	3.22 (4.13)

Figure in parenthesis is critical value; *** denotes rejection of hypothesis at 0.01 level, ** denotes rejection at 0.05 level

Table 8: Cointegrating Vector

	TGG	BCI
Coint. Vector	1.00	-1.76

Normalized ECM: $TGG = 1.76BCI$

There is one cointegrating vector, and BCI shows positive sign. This once again indicates a positive relationship between BCI and TGG.

Table 9: Adjustment Coefficients

Error Correction	D(TGG)	D(BCI)
CointEq1	-0.014 [-4.104]***	0.012 [1.29]
D(TGG(-1))	0.419 [2.439]**	0.917 [1.985]**
D(TGG(-2))	-0.267 [-1.569]	0.396 [0.86]
D(BCI(-1))	0.010 [0.159]	-0.083 [-0.465]
D(BCI(-2))	0.036 [0.59480]	0.146 [0.885]
R-squared	0.196	0.191
F-statistic	2.011	1.944

Figures in parentheses are t-values; *** denotes significance at 0.01 level; ** at 0.05 level; * at 0.10 level

This table gives the estimated coefficients of the variables in the ECM. We find that TGG has a statistically significant error correction coefficient with the correct sign. The adjustment coefficients indicate that

1.4% disequilibrium in TGG is corrected each year. BCI has no significant error correction. Further, we fail to find any significant t-values. However, on examining the individual ECM regressions, we find that TGG affects BCI. So, there must be causality from TGG to BCI.

After checking for cointegration and error correction between our two sets of variables, we now check for causality between finance and growth, by running tests of granger causality. This will confirm our findings from the earlier VECM.

Table 6: Granger Causality Tests

Null Hypothesis	Chi-sq	df	Prob.	F-statistic	Prob.
TBC does not cause TGG	12.56***	2	0.002	4.22**	0.012
TGG does not cause TBC	0.91	2	0.633	0.49	0.68
BCI does not cause TGG	0.36	2	0.835	0.49	0.69
TGG does not cause BCI	7.4**	2	0.025	2.03	0.129

*** denotes rejection of hypothesis at 0.01 level, ** denotes rejection at 0.05 level

The tests indicate that while TBC causes TGG, TGG doesn't cause TBC. This confirms the previous finding of unidirectional causality from TBC to TGG. However, we find that TGG causes BCI, but BCI doesn't cause TGG. This confirms the previous finding of unidirectional causality from TGG to BCI.

CONCLUSION

This study utilized forty years of Indian macro-economic data, from 1972 to 2012, to analyze the causal relationship between Bank credit and GDP growth. The results are interesting, indicating unidirectional causality from total bank credit to economic growth; and from economic growth to bank credit to industry. This leads us to believe that financial availability, in the form of bank credit, does indeed lead to economic growth. Once there is economic growth, there is a further positive effect on provision of bank credit, especially to the industry.

In an economy that is growing rapidly and where there is a thriving industry, there will be further need of credit, especially in the manufacturing and industrial sectors. This particular trend has been observed from our analysis. Another important point to remember is that during periods of economic growth, most industries are looking to expand operations, for which they borrow money. During periods of slow economic growth, if

banks and other financial institutions provide funds at attractive rates, it is bound to act as a catalyst to raise economic growth. Similarly, once adequate growth is achieved, it will lead to further expansion of the banking sector itself, and lead to a greater availability of industrial credit. There will also be a more diverse range of financial instruments and arrangements that crop up to meet the increased demand for credit.

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